

PQHS 471
Lecture 8: Resampling Methods
Cross-Validation, Bootstrap

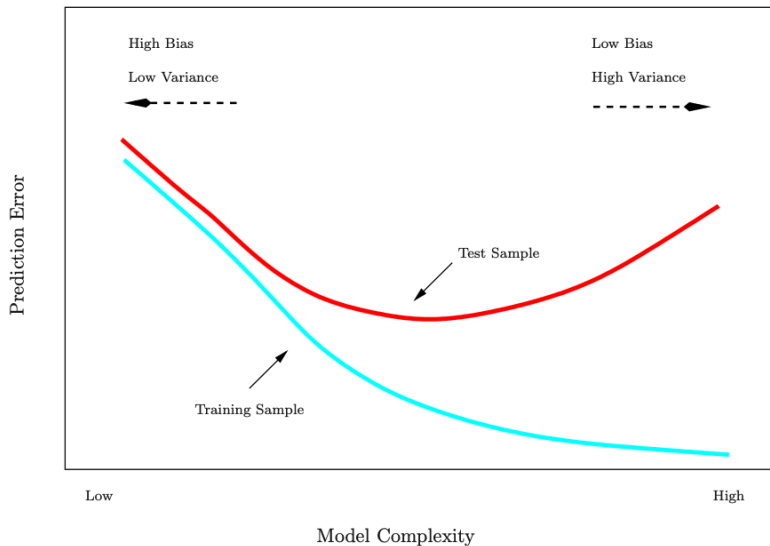
Cross-validation and the Bootstrap

- Two indispensable tools in modern machine learning and statistics.
- These methods refit a model of interest to samples formed from the training set, in order to obtain additional information about the fitted model.
- For example, they provide estimates of test-set prediction error, and the standard deviation and bias of our parameter estimates.

Training error vs test error

- Recall the distinction between the **test error** and the **training error**:
- The **test error** is the average error that results from using a statistical learning method to predict the response on a new observation, one that was not used in training the method.
- In contrast, the **training error** can be easily calculated by applying the statistical learning method to the observations used in its training.
- But the training error rate often is quite different from the test error rate, and in particular the former can **dramatically underestimate** the latter.

Training error vs test error



Prediction error estimate

- Best solution: a large designated test set. Often not available.
- Some methods make a mathematical adjustment to the training error rate. These include the Mallows's C_p , AIC and BIC.
- Here we instead consider a class of methods that estimate the test error by **holding out** a subset of the training observations from the fitting process, and then applying the statistical learning method to those held out observations.

Validation-set approach

- Here we randomly divide the available set of samples into two parts: a **training set** and a **validation** or **hold-out set**.
- The model is fit on the training set, and the fitted model is used to predict the responses for the observations in the validation set.
- The resulting validation-set error provides an estimate of the test error. This is typically assessed using MSE in the case of a quantitative response and misclassification rate in the case of a qualitative (discrete) response.

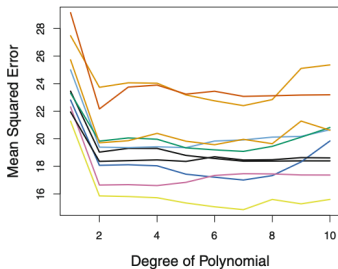
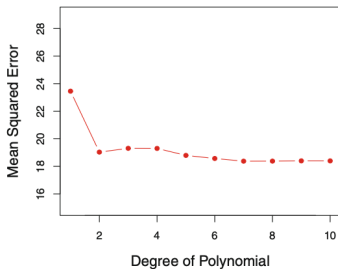
Validation-set approach



A random splitting into two halves: left part is training set, right part is validation set.

Example: automobile data

- Want to compare linear vs higher-order polynomial terms in a linear regression for predicting mpg.
- We randomly split the 392 observations into two sets, a training set containing 196 of the data points, and a validation set containing the remaining 196 observations.



Left panel shows single split; right panel shows 10 splits

Drawbacks of validation set approach

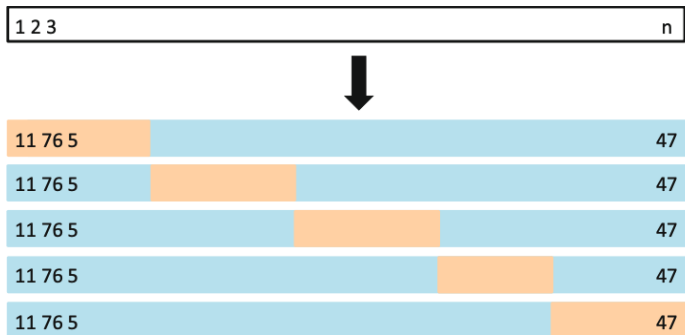
- Estimate of the test error can be highly variable, depending on precisely which observations are included in the training set and which observations are included in the validation set.
- Only a subset of the observations — those that are included in the training set rather than in the validation set — are used to fit the model.
- This suggests that the validation set error may tend to **overestimate** the test error for the model fit on the entire data set. **Why?**

K -fold Cross-validation

- **Widely used approach** for estimating test error.
- Estimates can be used to select best model, and to give an idea of the test error of the final chosen model.
- Idea is to randomly divide the data into K equal-sized parts. We leave out part k , fit the model to the other $K - 1$ parts (combined), and then obtain predictions for the left-out k th part.
- This is done in turn for each part $k = 1, 2, \dots, K$, and then the results are combined.

K -fold Cross-validation: example

Divide data into K roughly equal-sized parts ($K = 5$ here)



K -fold CV: details

- Let the K parts be C_1, C_2, \dots, C_K , where C_k denotes the indices of the observations in part k . There are n_k observations in part k : if N is a multiple of K , then $n_k = n/K$.
- Compute

$$CV_{(K)} = \sum_{k=1}^K \frac{n_k}{n} MSE_k$$

where $MSE_k = \sum_{i \in C_k} (y_i - \hat{y}_i)^2 / n_k$, and \hat{y}_i is the fit for observation i , obtained from the data with part k removed.

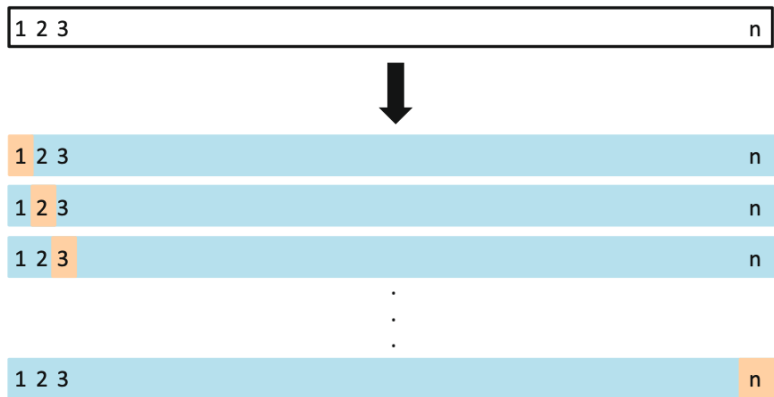
K -fold CV: details

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- Setting $K = n$ yields [leave-one out cross-validation](#) (LOOCV).



LOOCV: a nice special case

- With least-squares linear or polynomial regression, an amazing shortcut makes the cost of LOOCV the same as that of a single model fit! The following formula holds:

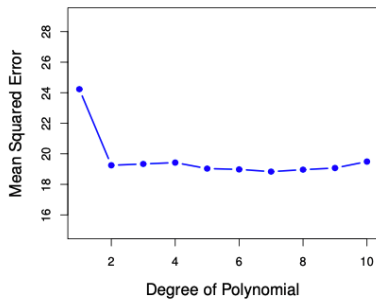
$$CV_{(n)} = \frac{1}{n} \sum_{i=1}^n \left(\frac{y_i - \hat{y}_i}{1 - h_i} \right)^2$$

where y_i is the i th fitted value from the original least squares fit, and h_i is the leverage (diagonal of the “hat” matrix) This is like the ordinary MSE, except the i th residual is divided by $1 - h_i$.

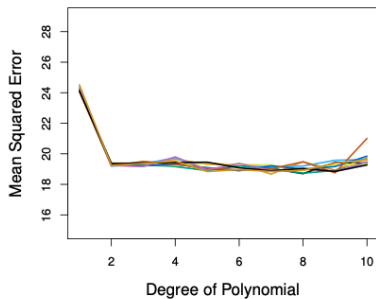
- LOOCV sometimes useful, but typically doesn't **shake up** the data enough. The estimates from each fold are highly correlated.

Auto data revisit

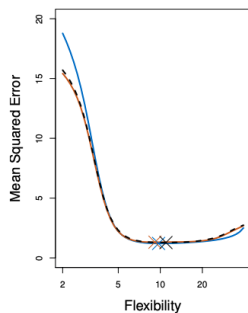
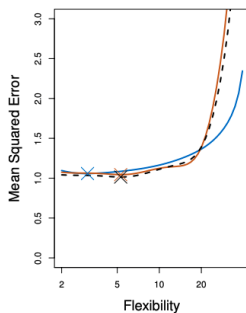
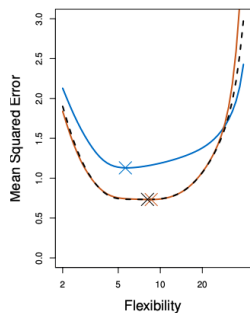
LOOCV



10-fold CV



True and est. test MSE in simulation



Blue: true test error in simulation.

Black dashed: LOOCV. Orange: 10-fold CV.

Considerations in cross-validation

- Since each training set is only $(K - 1)/K$ as big as the original training set, the estimates of prediction error will typically be biased upward. **Why?**
- This bias is minimized when $K = n$ (LOOCV), but this estimate has high variance.
- $K = 5$ or 10 provides a good compromise for this **bias-variance tradeoff**.

CV for classification problems

- Let the K parts be C_1, C_2, \dots, C_K , where C_k denotes the indices of the observations in part k . There are n_k observations in part k : if N is a multiple of K , then $n_k = n/K$.
- Compute:

$$CV_K = \sum_{k=1}^K \frac{n_k}{n} Err_k$$

where $Err_k = \sum_{i \in C_k} I(y_i \neq \hat{y}_i) / n_k$

Consider a simple classifier applied to some two-class data:

- 1 Starting with 5000 predictors and 50 samples, find the 100 predictors having the largest correlation with the class labels.
- 2 We then apply a classifier such as logistic regression, using only these 100 predictors.

How do we estimate the test set performance of this classifier?

Can we apply cross-validation in step 2, forgetting about step 1?

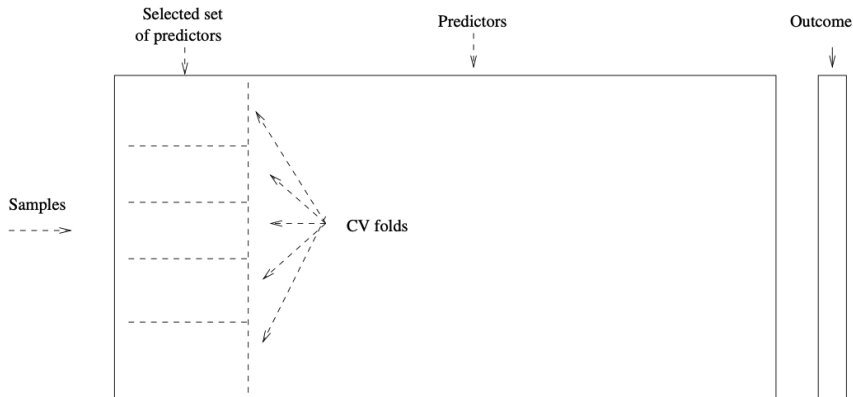
The answer is NO.

- This would ignore the fact that in Step 1, the procedure **has already seen the labels of the training data**, and made use of them. This is a form of training and must be included in the validation process.
- It is easy to simulate realistic data with the class labels independent of the outcome, so that true test error = 50%, but the CV error estimate that ignores Step 1 is zero! **Try to do this yourself**
- This is a common error in genomics research.

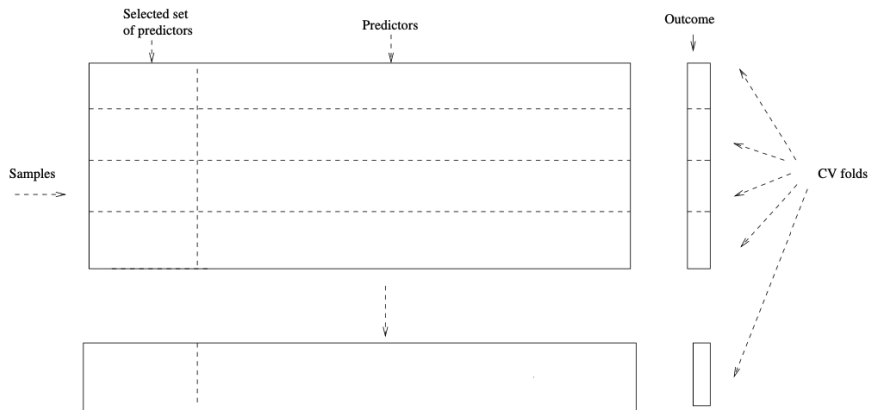
The Wrong and Right Way

- **Wrong:** Apply cross-validation in step 2.
- **Right:** Apply cross-validation to steps 1 and 2.

Wrong way



Right way



The Bootstrap

The Bootstrap

- The **bootstrap** is a flexible and powerful statistical tool that can be used to quantify the uncertainty associated with a given estimator or statistical learning method.
- For example, it can provide an estimate of the standard error of a coefficient, or a confidence interval for that coefficient.



- The use of the term bootstrap derives from the phrase **to pull oneself up by one's bootstraps**, widely thought to be based on one of the eighteenth century “The Surprising Adventures of Baron Munchausen” by Rudolph Erich Raspe:
The Baron had fallen to the bottom of a deep lake. Just when it looked like all was lost, he thought to pick himself up by his own bootstraps.

An example

- Suppose that we wish to invest a fixed sum of money in two financial assets that yield returns of X and Y , respectively, where X and Y are random quantities.
- We will invest a fraction α of our money in X , and will invest the remaining $1 - \alpha$ in Y .
- We wish to choose α to minimize the total risk, or variance, of our investment. In other words, we want to minimize $Var(\alpha X + (1 - \alpha)Y)$.
- One can show that the value that minimizes the risk is given by:

$$\alpha = \frac{\sigma_Y^2 - \sigma_{XY}}{\sigma_X^2 + \sigma_Y^2 - 2\sigma_{XY}}$$

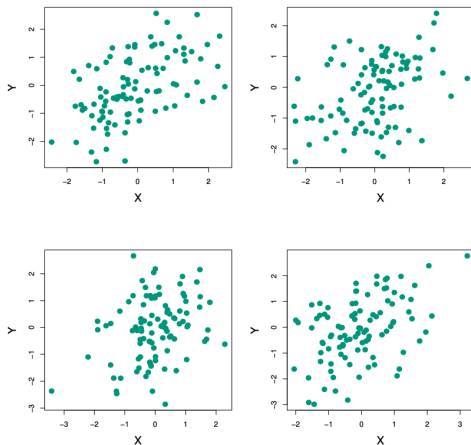
where $\sigma_X^2 = Var(X)$, $\sigma_Y^2 = Var(Y)$, and $\sigma_{XY} = Cov(X, Y)$.

An example

- But the values of σ_X^2 , σ_Y^2 and σ_{XY} are unknown.
- We can compute estimates for these quantities, $\hat{\sigma}_X^2$, $\hat{\sigma}_Y^2$ and $\hat{\sigma}_{XY}$, using a data set that contains measurements for X and Y .
- We can then estimate the value of α that minimizes the variance of our investment using

$$\hat{\alpha} = \frac{\hat{\sigma}_Y^2 - \hat{\sigma}_{XY}}{\hat{\sigma}_X^2 + \hat{\sigma}_Y^2 - 2\hat{\sigma}_{XY}}$$

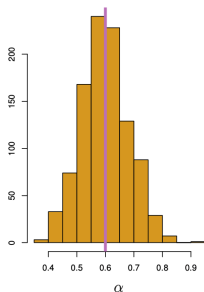
An example



Each panel displays 100 simulated returns for investments X and Y . From left to right and top to bottom, the resulting estimates for α are 0.576, 0.532, 0.657, and 0.651.

An example

- To estimate the standard deviation of $\hat{\alpha}$, we repeated the process of simulating 100 paired observations of X and Y , and estimating α 1,000 times.
- We thereby obtained 1,000 estimates for α , which we can call $\hat{\alpha}_1, \hat{\alpha}_2, \dots, \hat{\alpha}_{1000}$.
- For these simulations the parameters were set to $\sigma_X^2 = 1, \sigma_Y^2 = 1.25$, and $\sigma_{XY} = 0.5$, and so we know that the true value of α is 0.6 (indicated by the red line).



An example

- The mean over all 1,000 estimates for α is

$$\bar{a} = \frac{1}{1000} \sum_{r=1}^{1000} \hat{\alpha}_r = 0.5996$$

very close to $\alpha = 0.6$, and the standard deviation of the estimates is

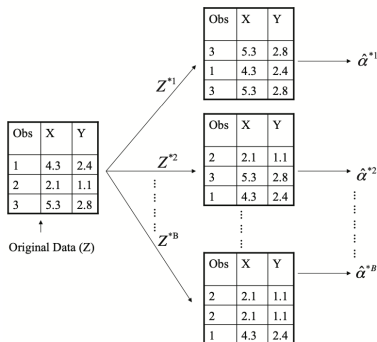
$$\sqrt{\frac{1}{1000 - 1} \sum_{r=1}^{1000} (\hat{\alpha}_r - \bar{a})^2} = 0.083$$

- This gives us a very good idea of the accuracy of $\hat{\alpha}$: $SE(\hat{\alpha}) \approx 0.083$.

Now back to the real world

- The procedure outlined above cannot be applied, because for real data we cannot generate new samples from the original population.
- However, the bootstrap approach allows us to use a computer to mimic the process of obtaining new data sets, so that we can estimate the variability of our estimate without generating additional samples.
- Rather than repeatedly obtaining independent data sets from the population, we instead obtain distinct data sets by repeatedly sampling observations from the original data set **with replacement**.
- Each of these “bootstrap data sets” is created by sampling **with replacement**, and is the **same size** as our original dataset. As a result some observations may appear more than once in a given bootstrap data set and some not at all.

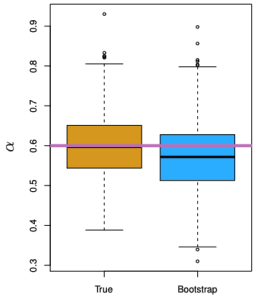
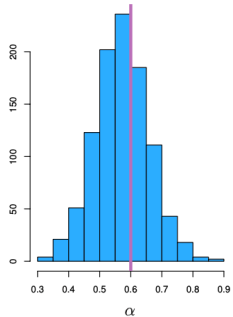
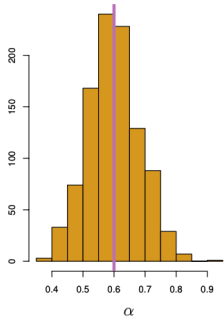
Example with just 3 observations



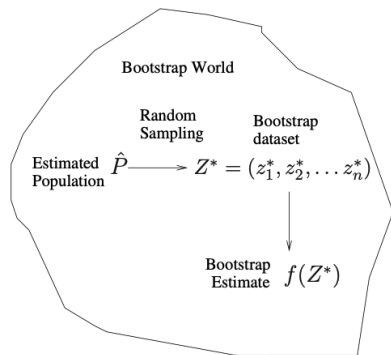
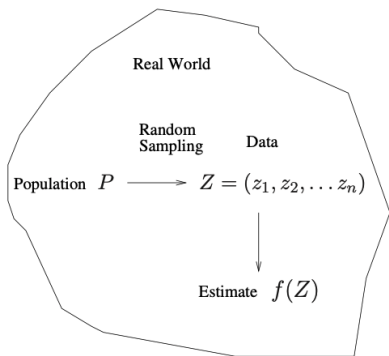
A graphical illustration of the bootstrap approach on a small sample containing $n = 3$ observations. Each bootstrap data set contains n observations, sampled with replacement from the original data set. Each bootstrap data set is used to obtain an estimate of α .

- Denoting the first bootstrap data set by Z^{*1} , we use Z^{*1} to produce a new bootstrap estimate for α , which we call α^{*1} .
- This procedure is repeated B times for some large value of B (say 100 or 1000), in order to produce B different bootstrap data sets, $Z^{*1}, Z^{*2}, \dots, Z^{*B}$ and B corresponding α estimates, $\alpha^{*1}, \alpha^{*2}, \dots, \alpha^{*B}$.
- We estimate the standard error of these bootstrap estimates using the formula:

$$\text{SE}_B(\hat{\alpha}) = \sqrt{\frac{1}{B-1} \sum_{r=1}^B (\hat{\alpha}^{*r} - \bar{\hat{\alpha}}^*)^2}$$



A general picture for the bootstrap



The bootstrap in general

- In more complex data situations, figuring out the appropriate way to generate bootstrap samples can require some thought.
- For example, if the data is a time series, we can't simply sample the observations with replacement.
- We can instead create blocks of consecutive observations, and sample those with replacements. Then we paste together sampled blocks to obtain a bootstrap dataset.

The bootstrap usage

- Primarily used to obtain standard errors of an estimate.
- Also provides approximate confidence intervals for a population parameter. For example, looking at the histogram in the middle panel of the Figure for α , the 5% and 95% quantiles of the 1000 values is (.43, .72).
- This represents an approximate 90% confidence interval for the true α .
- The above interval is called a **Bootstrap Percentile** confidence interval. It is the simplest method (among many approaches) for obtaining a confidence interval from the bootstrap.

- ISLR: chapter 5: 5.1 - 5.2